

Price Synchronization and Cost Pass-through in Multiproduct Firms: Evidence from Danish Producer Prices

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Inflation in a Changing Economic Environment ECB, Frankfurt, 23-24 September 2019

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Micro Price Stickiness & Monetary Neutrality

- <u>State-dependent</u> models can generate real effects similar to time-dependent price-setting when "selection" weak
 - If multiproduct firms synchronize their price changes (Midrigan 2011, Alvarez & Lippi 2014, Karadi & Reiff 2018)
- Micro evidence on actual price setting by <u>multiproduct</u> firms
 - Bhattarai & Schoenle (2014), Bonomo et al. (2019)
- <u>This paper</u>: Extensive (selection) and intensive margin of price adjustment in response to cost shocks in multiproduct firms
 - Micro data underlying Danish Producer Price Index (PPI)
 - Merge price and firm cost data

Main Results

- <u>Extensive margin</u> (whether to change prices):
 - Imperfect within & across firm synchronization of price changes
 - State-dependent pricing: Probability of changing prices affected by aggregate, industry and firm cost shocks
- <u>Intensive margin</u> (how much to change prices):
 - Despite state-dependence, small "selection" bias
- Cost-shocks pass-through heterogeneous
 - Less than complete (<<1) but immediate for import prices
 - Complete (≈1) but delayed for energy prices/oil supply shocks
 - Firms with 5+ products adjust less to import prices

Roadmap

1. Data and descriptive statistics

- 2. Empirical approach
- 3. Results

Data: Prices and Firms

- Monthly goods prices for <u>Danish PPI</u> covering 70+% total sales of industrial production, 1993-2017
 - 3500 monthly prices for domestic and export transactions
 - 2900 monthly imported input prices
 - Median duration of price reporting: 115 months
 - 1140 firms (not representative sample)
- Merge with firm-level cost data:
 - Accounting data: Annual cost shares, 1994-2016
 - VAT filings: Monthly/quarterly sales & input purchases, 2001-2017
 - Labor costs: Monthly wage bill and hours worked, 2007-2017

Summary Statistics for Prices and Firms

	All	1	1-3	3-5	5-7	7+
No. of firms	1140	146	548	231	128	87
Median employment (FTE)	155.0	42.6	65.5	138.6	148.3	483.1
Median employment per goo	d 33.2	42.6	26.2	34.8	25.5	48.5
Median age	26.0	25.0	25.0	28.0	24.0	29.0
Share of total prices	100.0	2.1	20.2	22.8	16.7	38.1
Median no. of products	5.0	1.0	3.0	4.0	5.8	11.5

Note: Summary statistics on distribution of firms and prices across distinct bins of single- and multiproduct firms.

Frequency of Price Increases & Decreases



- Median frequency $\approx 10\% => 8$ months price duration
- Seasonal patterns in January, April, July, October (Nakamura & Steinsson, 2008 and IPN)

Average Price Changes



Average absolute size of Δp ≈7%

• "Sales" only 0.3% (3.8% of decreases but smaller drop)

Size Distribution of Non-zero Price Changes



- Large mass of small Δp (Alvarez, Le Bihan & Lippi 2016)
- Distribution very similar across firms' # of goods

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Price Adjustment to Cost Shocks

- Under <u>flex prices</u>, $\Delta p = \Delta markup + \Delta (marginal cost)$
 - <u>Structural</u> cost pass-through: $\partial \Delta p / \partial (\Delta (marginal cost))$
 - Pass-through < 1 when markup absorbs cost increases

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 - <u>Structural</u> cost pass-through: $\partial \Delta p / \partial (\Delta (marginal cost))$
 - Pass-through < 1 when markup absorbs cost increases
- Under <u>sticky prices</u> adjustment through <u>extensive</u> and <u>intensive</u> margin
 - With infrequent adjustment downward bias if $\Delta p=0$ included
 - Looking only at ∆p≠0 leads to selection bias with statedependent extensive margin
 - Both margins matter for inflationary effects of cost shocks: OLS of Δp on costs gives overall elasticity, but not structural

Estimating Extensive and Intensive Margins

- To correct for selection bias use two-stage "Heckit" approach
- Jointly model both margins following Bourguignon et al. (2007): 1st stage multinomial logit selection

 $\Delta p_{i,j,t}^* = \beta^1 Z_{i,j,t} + \eta_{i,j,t} \rightarrow \qquad \Pr(r_{i,j,t} = -1, 0, 1) = \Phi(\beta Z_{i,j,t})$ $\Delta p_{i,j,t,r=1} = \beta^2 X_{i,j,t,r=1} + u_1$

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• u correlated with η , assume function of estimated probabilities:

$$\Delta p_{i,j,t,r=1} = \beta^2 X_{i,j,t,r=1} + u_1$$

= $\beta^2 X_{i,j,t,r=1} + \gamma_1 m(\Pr_1) + \sum_{r=(-1,0)} \gamma_r \left(m(\Pr_r) \frac{\Pr_r}{(\Pr_r - 1)} \right) + \omega_1$

• Non-parametric identification, X subset of Z

Estimating Extensive and Intensive Margin

- Identification by exploiting price synchronization in multiproduct firms and frequency seasonality (all excluded from 2nd stage X):
 - Share of same and opposite-signed price changes within firm and sector (excluding i-th price)
 - Month fixed effects
 - Also standard deviation of specific price in last 5 years

Dependent Variable in Selection Step



Probabilities of cumulative price changes

• 40% of cumulated $\Delta p=0$ even after 12 months

Estimating Intensive Margin

- Marginal cost = input prices weighted by shares in total variable costs (Amiti et al. 2018)
- Local projections to estimate t+h cost pass-through:

 $\Delta p_{i,j,t+h} = \alpha + \beta^E \varphi^E_{j,t-1} \Delta p^E_t + \beta^I \varphi^I_{j,t-1} \Delta p^M_{j,t} + \partial X_{i,j,t} + u_{i,j,t+h}$

- Firm-level cost shocks:
 - Δenergy price x cost share of energy(t-1) average 1.7%
 - Δimport prices x cost share of imports(t-1) average 28%

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- Further firm controls (in X): change in firm hourly wages, change in total variable costs, change in total sales; firm size and #products
 - Also proxy for markup with prices of competitors (2 digit level)
 - Aggregate controls: CPI, NEER (and sector FE)

Energy and Imported Input Prices

Danish Energy Price Index

Average Imported Input Price



- Both approximately random walks (shocks *i.i.d.*)
 => Pass-through similar across horizons
- But small common component in energy prices

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Results #1: Determinants of Extensive Margin

- Within and across firms synchronization of price changes
 - Within-firm share of price changes of opposite sign
 - Increasing in number of products of firm

- Evidence of state-dependent pricing -- Probability of changing prices affected by firm, industry and aggregate shocks
 - 1% increase in CPI raises probability $\Delta p>0$ by 0.56%
 - 1% increase in import prices raises probability $\Delta p>0$ by 0.28%

Extensive Margin: Synchronization

	All		
Marg. effect on probability of price increase		Marg. effect on probability of price decrease	
Fraction of pos. price changes in firm	6.34***	Fraction of pos. price changes in firm	2.28***
	(0.04)		(0.04)
Fraction of neg. price changes in firm	2.74***	Fraction of neg. price changes in firm	4.09***
	(0.04)		(0.03)
Fraction of pos. price changes in industry	0.080	Fraction of pos. price changes in industry	-0.25***
	(0.06)		(0.06)
Fraction of neg. price changes in industry	-0.202**	Fraction of neg. price changes in industry	-0.073
	(0.06)		(0.06)

Imperfect synchronization within and across firms

Extensive Margin: Synchronization

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Marg. effect on probability of price increase		Marg. effect on probability of price decrease	
Fraction of pos. price changes in firm	6.34***	Fraction of pos. price changes in firm	2.28***
	(0.04)		(0.04)
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	(0.06)		(0.06)
Fraction of neg. price changes in industry	-0.202**	Fraction of neg. price changes in industry	-0.073
	(0.06)		(0.06)

Imperfect synchronization within and across firms

Extensive Margin: Changing # of Products

	All	1-5	5+
Marg. effect on probability of price increase			
Fraction of pos. price changes in firm	6.34***	5.27***	7.83***
	(0.04)	(0.04)	(0.06)
Fraction of neg. price changes in firm	2.74 * * *	2.39***	2.87***
	(0.04)	(0.05)	(0.07)
Fraction of pos. price changes in industry	0.080	0.333 * *	0.037
	(0.06)	(0.11)	(0.08)
Fraction of neg. price changes in industry	-0.202**	-0.43***	-0.104
	(0.06)	(0.13)	(0.08)

Within firm synchronization increasing, but effect small Across firm synchronization decreasing with # of products

Extensive Margin: State Dependence

	All		X /
Marg. effect on probability of price increase		Marg. effect on probability of price decrease	
Fraction of pos. price changes in firm	6.34***	Fraction of pos. price changes in firm	2.28***
	(0.04)		(0.04)
Fraction of neg. price changes in firm	2.74***	Fraction of neg. price changes in firm	4.09***
	(0.04)		(0.03)
Fraction of pos. price changes in industry	0.080	Fraction of pos. price changes in industry	-0.25***
	(0.06)		(0.06)
Fraction of neg. price changes in industry	-0.202**	Fraction of neg. price changes in industry	-0.073
	(0.06)		(0.06)
Avg. price change in industry, excl. firm	0.14***	Avg. price change in industry, excl. firm	-0.15***
	(0.03)		(0.03)
Energy price change x lagged energy cost share	-0.371	Energy price change x lagged energy cost share	-0.172
	(0.38)		(0.34)
Import price change x lagged import cost share	0.28***	Import price change x lagged import cost share	-0.29***
	(0.04)		(0.04)
CPI, log difference	0.557*	CPI, log difference	-1.00***
	(0.28)		(0.27)

State-dependence wrt industry, firm and aggregate variables (Energy price significant after 3 months)

Results #2: Intensive Margin

- Despite state-dependence, little evidence of selection bias
 - Bias correction terms significant but small
- Cost-shocks pass-through heterogeneous
 - Less than complete (<<1) but immediate for import prices
 - Complete (≈1) but gradual for energy prices/oil supply shocks
 - Firms with 5+ products adjust less to import prices

Heterogeneous Cost Pass-Through: OLS

Energy cost pass-through, incl. zero changes

Import price pass-through, incl. zero changes



Heterogeneous Cost Pass-Through

(a) Energy cost pass-through

(b) Import price pass-through



Dynamics robust to excluding lags of energy and import costs (Caveat: OLS (HAC) standard errors)

Heterogeneous Cost Pass-Through: Oil Shock

(a) Oil price surprise pass-through

(b) Import price pass-through



Negative oil supply shock (i.i.d.), Baumeister & Hamilton (2019)

Heterogeneous Cost Pass-Through: # Goods

Firms with **#** Goods ≤5

Firms with # Goods>5



Import price pass-through, excl. zero changes

Import price pass-through, excl. zero changes

Heterogeneity in markup adjustment (despite competitors prices)? But very similar responses to energy prices/oil shocks

To conclude

- Multiproduct firms' extensive and intensive margin of price adjustment
- <u>Synchronization and state-dependence</u> in extensive margin:
 - Price change probability increasing with fraction of other prices changing, higher with more products
 - Affected by firm, industry and aggregate shocks
- <u>Cost and firm heterogeneity</u> in intensive margin:
 - Adjustment to energy/oil shocks larger than to import prices
 - Firms with more products adjust by less to import prices
- Still preliminary, next up: markups, non-linearities, domestic and export prices, monetary shocks,...

Price Setting by Multiproduct Firms

- **Does simple** Δp pattern depend on # of goods by firm?
 - Alvarez-Lippi (2014): Δp frequency (N) and size depend on n

$$E\left[N\left(\Delta p_{i}\right)\right] = \frac{n\sigma^{2}}{\overline{y}} = \frac{n\sigma^{2}}{\sqrt{2\left(n+2\right)\frac{\sigma^{2}\psi}{B}}} \uparrow \text{ in } n$$
$$E\left\|\Delta p_{i}\right\| = \frac{\sqrt{\overline{y}}}{\frac{n-1}{2}Beta\left(\frac{n-1}{2},1/2\right)} = \frac{\left(2\left(n+2\right)\frac{\sigma^{2}\psi}{B}\right)^{1/4}}{\frac{n-1}{2}Beta\left(\frac{n-1}{2},1/2\right)} \downarrow \text{ in } n$$

- Consistent evidence in Bhattarai & Schoenle (2014) for US PPI
 - We find little relation between # goods and Δp frequency & size

What We Do

- Empirical analysis of relative price adjustment in a <u>currency</u> area during housing bust in the US Great Recession
- Build new dataset of <u>regional (MSA) consumption prices</u>, combined with data on sectoral costs and activity:
 - Based on BLS CPI Research database for 73 MSAs
 - Public CPI data for 27 largest MSAs
 - BLS micro data for smaller MSAs
- <u>Decomposition</u> of real exchange rate adjustment between goods and services (ex-rents) across US MSAs
- Empirical approach: (sectoral) prices, employment, wages,..., regressed on local house prices as a measure of demand – both OLS and IV

Main Results: Missing Internal Devaluation

- Little real exchange rate adjustment to asymmetric housing bust within US "currency area"
 - Evidence that the relative price of (non-tradable) services does not fall more than relative price of (tradable) goods
 - <u>Relative price of goods</u> insensitive to house prices
 - <u>Relative price of services</u> negatively related to house prices
- <u>Heterogeneity</u> in price adjustment across related sectors
 - Price of Food at home <u>falls</u> with bust in house prices
 Stroebl & Vavra 2018 (S&V18)
 - Price of Food away from home significantly increases

Selected Literature

- Large literature on the link between regional prices and wages, and regional business cycles (Blanchard & Katz 1992)
- Recent contributions:
 - Local employment and demand effects of house price shocks: Mian, Rao & Sufi (2013), Mian & Sufi (2014), Aladangady (2016), Kaplan, Mitman & Violante (2016)
 - Local prices/markups using store-level scanner data: Coibion et al. (2014), Stroebl & Vavra (2018), Beraja, Hurst & Ospina (2016), Kaplan et al. (2016), Anderson, Rebelo & Wong (2018)
 - <u>Regional Phillips Curves</u>: Fitzgerald & Nicolini (2014)
 - <u>Regional and sectoral adjustment</u> in currency areas: Philippon & Midrigan (2014), Beraja et al. (2016), Martin & Philippon (2017), Galí & Monacelli (2018) – and Nakamura & Steinsson (2014) on fiscal shocks,...

Stylized Adjustment Mechanism

- Demand for locally produced tradables depends on overall aggregate demand in currency area
- Non-tradable (services) demand depends on local demand *D* and relative prices:

$$logY_{l,N} = -\epsilon_N \cdot (1 - \alpha_N) \cdot (p_{l,N} - p_{l,T}) + logD^l$$

• A <u>fall</u> in *D* can be cushioned by a <u>fall in relative prices</u> $(p_{l,N} - p_{l,T})$ and in local marginal cost (function of labor demand):

$$p_{l,N} = log\mu_{l,N} + \gamma_N logW_l + (1 - \gamma_N) logl_{l,N}$$
$$p_{l,T} = (log\mu_{l,T}) + \gamma_T logW_l + (1 - \gamma_T) log(L_l - l_{l,N})$$

• Ceteris paribus, fall in relative price will result in *rer* depreciation: $rer_l \downarrow = (p_{l,T} - p_T) + \alpha_{NT} \cdot [(p_{l,NT} - p_{l,T}) \downarrow - (p_{NT} - p_T)]$ $rer_l \downarrow = q_{l,G} + \alpha_S \cdot [q_{l,S} \downarrow - q_{l,G}]$

Procyclical Markup Adjustment?

- S&V18 shows that positive elasticity of scanner grocery prices is due to procyclical markup adjustment
 - Similar finding with BLS index no "trading down" (Jaimovich, Rebelo & Wong 2108)
- Evidence consistent with non-homothetic preferences (Bertoletti & Etro 2017):

$$\mu^l = 1 + \frac{y^l}{\gamma * costs}$$

- Markup then falls with local income/wealth if costs acyclical (Anderson, Rebelo & Wong 2018)
- But then why are markups countercyclical in restaurants, and services in general, beyond movements in costs?

Countercyclical Markup Adjustment?

• Static oligopoly model after Atkeson & Burstein (2008): Markup function of market shares and within/across sectors elasticities (θ_j, φ)

$$\mu_j = \frac{\epsilon_j}{\epsilon_j - 1}, \epsilon_j = \theta_j - (\theta_j - \varphi) \cdot share_j$$

- Markup rises (elasticity falls) when market share rises (# of firms falls) if (θ_j φ) > 0
- Evidence with other measures of markups than labor shares?
 Bils, Klenow and Malin (2018): try share of intermediate inputs